

Multivariate Data

- Previous methods on sample: $\mathcal{X} = \{x^t\}_{t=1}^N$ $\mathcal{X} = \{x^t, r^t\}_{t=1}^N$
- Several measurements generate observation vector. Sample as a data matrix

$$\mathbf{X} = \begin{bmatrix} X_1^1 & X_2^1 & \cdots & X_d^1 \\ X_1^2 & X_2^2 & \cdots & X_d^2 \\ \vdots & & & \\ X_1^N & X_2^N & \cdots & X_d^N \end{bmatrix}$$

- Nxd matrix
- d columns: d variables
- N rows: N instances
 - Independent and identically distributed (i.i.d)

Multivariate Data – Parameter Estimation

Mean vector

$$E[\mathbf{x}] = \boldsymbol{\mu} = [\mu_1, \dots, \mu_d]^T$$

The variance of X_i

$$\sigma_i^2$$

The covariance of two variables X_i and X_j

$$\sigma_{ij} \equiv \text{Cov}(X_i, X_j) = E[(X_i - \mu_i)(X_j - \mu_j)]$$

Covariance matrix

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1d} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2d} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{d1} & \sigma_{d2} & \cdots & \sigma_d^2 \end{bmatrix}$$

$$\Sigma \equiv \text{Cov}(\mathbf{X}) = E[(\mathbf{X} - \boldsymbol{\mu})(\mathbf{X} - \boldsymbol{\mu})^T]$$

Linear
dependence:
> 0
< 0
= 0

Multivariate Data – Parameter Estimation

Covariance matrix

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1d} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2d} \\ \vdots & & & \\ \sigma_{d1} & \sigma_{d2} & \cdots & \sigma_d^2 \end{bmatrix}$$

$$\Sigma \equiv \text{Cov}(X) = E[(X - \mu)(X - \mu)^T]$$

Correlation

$$\text{Corr}(X_i, X_j) \equiv \rho_{ij} = \frac{\sigma_{ij}}{\sigma_i \sigma_j}$$

Two variables independent $\rightarrow$ covariance = correlation = 0

Covariance = 0 $\rightarrow$ two variables independent

Multivariate Data – Parameter Estimation

Sample mean $\mathbf{m} = \frac{\sum_{t=1}^N \mathbf{x}^t}{N}$ with $m_i = \frac{\sum_{t=1}^N x_i^t}{N}, i = 1, \dots, d$

Sample covariance matrix

$$s_i^2 = \frac{\sum_{t=1}^N (x_i^t - m_i)^2}{N}$$
$$s_{ij} = \frac{\sum_{t=1}^N (x_i^t - m_i)(x_j^t - m_j)}{N}$$

Multivariate Data – Estimating missing values

- Fill-in missing entries by estimating them: **imputation**
 - Mean imputation
 - Imputation by regression

	sex	race	educ_r	r_age	earnings	police
[91,]	1	3	3	31	NA	0
[92,]	2	1	2	37	135.00	1
[93,]	2	3	2	40	NA	1
[94,]	1	1	3	42	3.00	1
[95,]	1	3	1	24	0.00	NA

Multivariate Data – Normal distribution

$$\mathbf{x} \sim \mathcal{N}_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

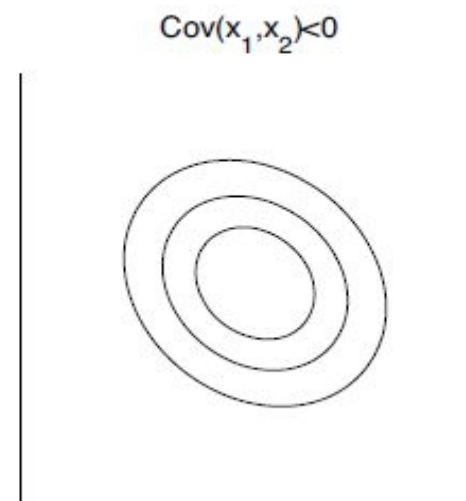
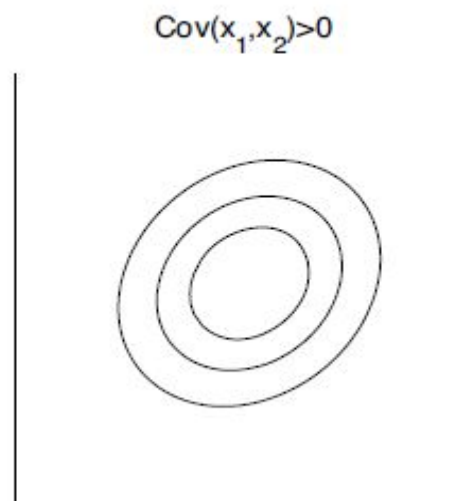
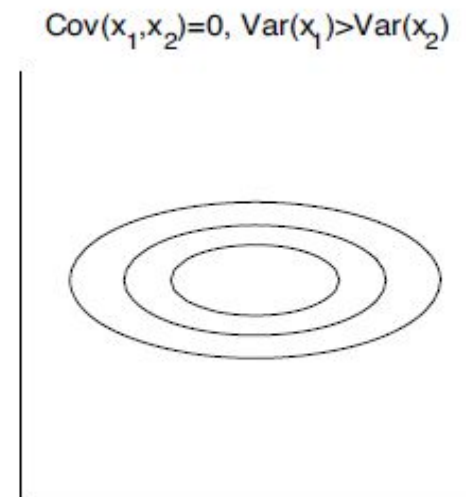
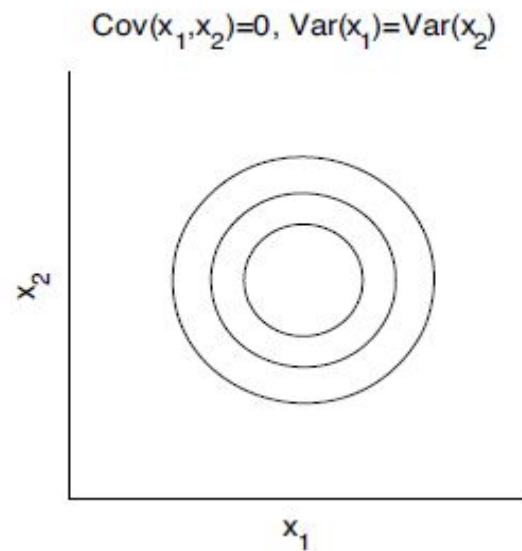
$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right]$$

Univariate case: $\frac{(x - \mu)^2}{\sigma^2} = (x - \mu)(\sigma^2)^{-1}(x - \mu)$



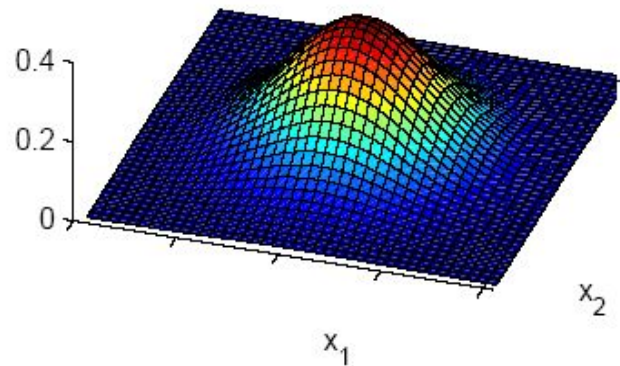
Mahalanobis distance
(standardizes all variables
to unit variance and eliminate
correlations)

Multivariate Data – Normal distribution

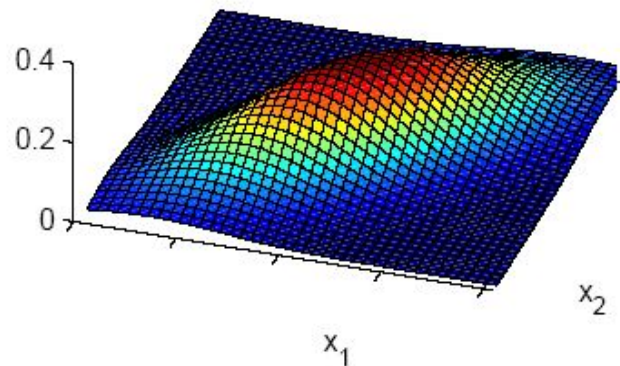


Multivariate Data – Normal distribution

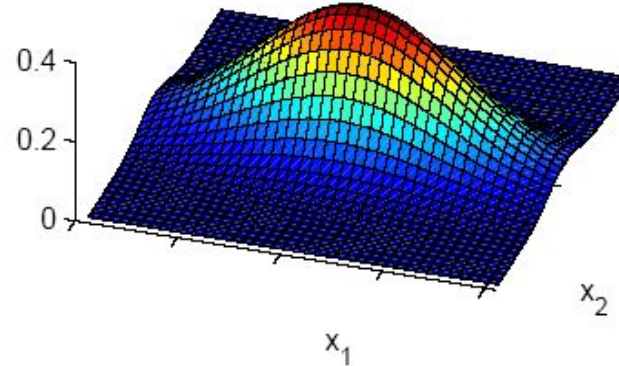
$\text{Cov}(x_1, x_2) = 0, \text{Var}(x_1) = \text{Var}(x_2)$



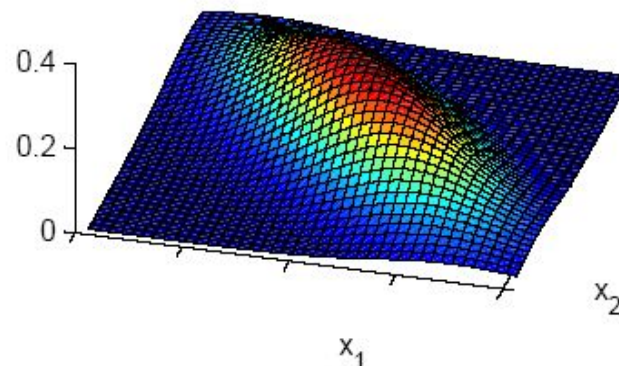
$\text{Cov}(x_1, x_2) > 0$



$\text{Cov}(x_1, x_2) = 0, \text{Var}(x_1) > \text{Var}(x_2)$



$\text{Cov}(x_1, x_2) < 0$



Multivariate Data – Classification

$$p(\mathbf{x}|C_i) = \frac{1}{(2\pi)^{d/2}|\Sigma_i|^{1/2}} \exp \left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_i)^T \Sigma_i^{-1}(\mathbf{x} - \boldsymbol{\mu}_i) \right]$$

$$g_i(\mathbf{x}) = -\frac{d}{2} \log 2\pi - \frac{1}{2} \log |\Sigma_i| - \frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_i)^T \Sigma_i^{-1}(\mathbf{x} - \boldsymbol{\mu}_i) + \log P(C_i)$$

Using maximum likelihood:

$$\begin{aligned}\hat{P}(C_i) &= \frac{\sum_t r_i^t}{N} \\ \mathbf{m}_i &= \frac{\sum_t r_i^t \mathbf{x}^t}{\sum_t r_i^t} \\ \mathbf{S}_i &= \frac{\sum_t r_i^t (\mathbf{x}^t - \mathbf{m}_i)(\mathbf{x}^t - \mathbf{m}_i)^T}{\sum_t r_i^t}\end{aligned}$$

Multivariate Data – Classification

- Use common covariance matrix for each class
 - Assume all off-diagonals of covariance matrix be 0
 - Product of individual univariate densities...
-
- All variances equal – Mahalanobis distance to Euclidean distance
 - All priors are equal

Multivariate Data – Classification

- Use common covariance matrix for each class
- Assume all off-diagonals of covariance matrix be 0

Naive Bayes' classifier

$$g_i(\mathbf{x}) = -\frac{1}{2} \sum_{j=1}^d \left(\frac{x_j^t - m_{ij}}{s_j} \right)^2 + \log \hat{P}(C_i)$$

- All variances equal – Mahalanobis distance to Euclidean distance
- All priors are equal

Nearest mean classifier

$$g_i(\mathbf{x}) = -\|\mathbf{x} - \mathbf{m}_i\|^2$$

Multivariate Data – Linear Regression

Minimize sum of squared errors:

$$E(w_0, w_1, \dots, w_d | \mathcal{X}) = \frac{1}{2} \sum_t (r^t - w_0 - w_1 x_1^t - w_2 x_2^t - \dots - w_d x_d^t)^2$$

Take derivatives:

$$\begin{aligned} \sum_t r^t &= Nw_0 + w_1 \sum_t x_1^t + w_2 \sum_t x_2^t + \dots + w_d \sum_t x_d^t \\ \sum_t x_1^t r^t &= w_0 \sum_t x_1^t + w_1 \sum_t (x_1^t)^2 + w_2 \sum_t x_1^t x_2^t + \dots + w_d \sum_t x_1^t x_d^t \\ \sum_t x_2^t r^t &= w_0 \sum_t x_2^t + w_1 \sum_t x_1^t x_2^t + w_2 \sum_t (x_2^t)^2 + \dots + w_d \sum_t x_2^t x_d^t \\ &\vdots \\ \sum_t x_d^t r^t &= w_0 \sum_t x_d^t + w_1 \sum_t x_d^t x_1^t + w_2 \sum_t x_d^t x_2^t + \dots + w_d \sum_t (x_d^t)^2 \end{aligned}$$

Linear regression, higher order polynomial

$$\mathbf{A}\mathbf{w} = \mathbf{y}$$

$$\mathbf{A} = \begin{bmatrix} N & \sum_t x^t & \sum_t (x^t)^2 & \dots & \sum_t (x^t)^k \\ \sum_t x^t & \sum_t (x^t)^2 & \sum_t (x^t)^3 & \dots & \sum_t (x^t)^{k+1} \\ \vdots & & & & \\ \sum_t (x^t)^k & \sum_t (x^t)^{k+1} & \sum_t (x^t)^{k+2} & \dots & \sum_t (x^t)^{2k} \end{bmatrix}$$
$$\mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ w_2 \\ \vdots \\ w_k \end{bmatrix}, \mathbf{y} = \begin{bmatrix} \sum_t r^t \\ \sum_t r^t x^t \\ \sum_t r^t (x^t)^2 \\ \vdots \\ \sum_t r^t (x^t)^k \end{bmatrix}$$

We can write $\mathbf{A} = \mathbf{D}^T \mathbf{D}$ and $\mathbf{y} = \mathbf{D}^T \mathbf{r}$ where

$$\mathbf{D} = \begin{bmatrix} 1 & x^1 & (x^1)^2 & \dots & (x^1)^k \\ 1 & x^2 & (x^2)^2 & \dots & (x^2)^k \\ \vdots & & & & \\ 1 & x^N & (x^N)^2 & \dots & (x^N)^k \end{bmatrix}, \mathbf{r} = \begin{bmatrix} r^1 \\ r^2 \\ \vdots \\ r^N \end{bmatrix}$$

and we can then solve for the parameters as

$$\mathbf{w} = (\mathbf{D}^T \mathbf{D})^{-1} \mathbf{D}^T \mathbf{r}$$

Multivariate Data – Linear Regression

$$\mathbf{X} = \begin{bmatrix} 1 & x_1^1 & x_2^1 & \cdots & x_d^1 \\ 1 & x_1^2 & x_2^2 & \cdots & x_d^2 \\ \vdots & & & & \\ 1 & x_1^N & x_2^N & \cdots & x_d^N \end{bmatrix}, \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_d \end{bmatrix}, \mathbf{r} = \begin{bmatrix} r^1 \\ r^2 \\ \vdots \\ r^N \end{bmatrix}$$

Then the normal equations can be written as

$$\mathbf{X}^T \mathbf{X} \mathbf{w} = \mathbf{X}^T \mathbf{r}$$

and we can solve for the parameters as

$$\mathbf{w} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{r}$$